

HG  
Nov. 11

## ECON 4130 H11

### Extra exercises for no-seminar week 45

(Solutions will be put on the net at the end of the week)

#### Exercise 1

*This exercise is based on the Exam 2004H -“utsatt prøve”, slightly extended and adapted to fit the present curriculum.*

The random variable (rv.),  $Y$ , has a log-normal distribution with parameters,  $\mu$  and  $\sigma^2$ , if the density function (pdf) is given by

$$f(y) = \begin{cases} \frac{1}{\sqrt{2\pi\sigma}} \cdot \frac{1}{y} \cdot e^{-\frac{1}{2\sigma^2}[\ln(y)-\mu]^2} & \text{for } y > 0 \\ 0 & \text{otherwise} \end{cases}$$

This is a right skewed distribution with a pdf somewhat similar to the pdf of a gamma distribution. It is sometimes used to model income distributions.

- A. Show that, if  $Y$  is log-normal  $(\mu, \sigma^2)$  then  $X = \ln(Y)$  is normally distributed with expectation,  $\mu$  and variance,  $\sigma^2$  (i.e.,  $N(\mu, \sigma^2)$ ).
- B. Explain how the moment generating function (mgf) for  $X$ , can be utilized to show that

$$E(Y^k) = e^{k\mu + k^2 \frac{\sigma^2}{2}} \quad \text{for } k = 1, 2, 3, \dots$$

- C. The *variation coefficient* of a non-negative rv.,  $Z$ , denoted by  $VC(Z)$ , is defined as

$$VC(Z) = \frac{\sqrt{\text{Var}(Z)}}{E(Z)}$$

The variation coefficient is a measure of variation. If  $Z$  is the income of a person randomly chosen from a population of income earners,  $VC(Z)$  is sometimes taken as a measure of income inequality for the population in question.

- (i) Show that the VC is invariant for scale transformations, i.e., show that  $VC(cZ) = VC(Z)$  for any constant  $c > 0$ .
- (ii) Let  $Y$  be log-normal  $(\mu, \sigma^2)$ . Show that the variation coefficient, that we will denote by  $\gamma$ , is  $\gamma = VC(Y) = \sqrt{e^{\sigma^2} - 1}$
- (iii) Let  $Y$  be gamma distributed,  $(\alpha, \lambda)$ , where  $\alpha$  is the shape parameter and  $\lambda$  the scale parameter. Show that the variation coefficient is equal to  $1/\sqrt{\alpha}$  and, hence, independent of the scale  $\lambda$ .

**D.** Let  $Z_1, Z_2, \dots, Z_n$  be *iid* and non-negative rv's with expectation,  $E(Z_i) = \eta$ , and variance,  $\text{Var}(Z_i) = \tau^2$ . Otherwise we don't know anything about the common distribution of the  $Z_i$ 's. Propose a consistent estimator for the VC in this case, and explain why it is consistent.

**E.** Let  $Y_1, Y_2, \dots, Y_n$  be *iid* and log-normally distributed  $(\mu, \sigma^2)$ . Show that the maximum likelihood estimators (MLE's) for  $\mu$  and  $\sigma^2$  are given by

$\hat{\mu} = \overline{\ln(Y)} = \frac{1}{n} \sum_{i=1}^n \ln(Y_i)$  and  $\hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n [\ln(Y_i) - \overline{\ln(Y)}]^2$  respectively. [**Hint:** Compare the log likelihood function with the log likelihood for a normal sample, i.e., study example B in Rice section 8.5 ]

What is the MLE for the variation coefficient,  $\gamma = VC(Y_i)$ ?

**F.** Derive the moment estimators (MME's) for  $\mu, \sigma^2$ , and  $\gamma$ , based on  $Y_1, Y_2, \dots, Y_n$  in question E.

**G.** We have a sample of  $n = 121$  yearly incomes drawn from a population of women (Norway 1998) that is relatively homogenous with regard to the time spent at paid work. Let  $Y_i$  denote the income of woman  $i$  in the sample. As before we assume that  $Y_1, Y_2, \dots, Y_n$  is *iid* and log-normally distributed  $(\mu, \sigma^2)$ . Calculate both the MLE- and MME estimates of the population VC,  $\gamma$ , based on the summary data in the table

<i>Statistic</i>	<i>Data</i>
$n$ (sample size)	121
$\frac{1}{n} \sum_{i=1}^n Y_i$ (NOK)	202 799
$\frac{1}{n} \sum_{i=1}^n Y_i^2$	46 597 545 146
$\frac{1}{n} \sum_{i=1}^n \ln(Y_i)$	12.15916
$\frac{1}{n} \sum_{i=1}^n [\ln(Y_i)]^2$	147.96481

**H.** It can be shown that the MLE,  $\hat{\sigma}^2$ , is asymptotically normally distributed in the sense  $\sqrt{n}(\hat{\sigma}^2 - \sigma^2) \xrightarrow[n \rightarrow \infty]{D} N(0, 2\sigma^4)$  (You do not need to show this here.) Use this to develop an asymptotic 95% confidence interval (CI) for  $\gamma$  based on a corresponding CI for  $\sigma^2$ . Calculate the interval.

**I.** A well known fact is that if  $X_1, X_2, \dots, X_n$  are *iid* and normally distributed,

$X_i \sim N(\mu, \sigma^2)$ , then  $\frac{\sum_{i=1}^n (X_i - \bar{X})^2}{\sigma^2}$  is (exactly)  $\chi^2$  distributed with  $n-1$  degrees of freedom. Use this to find an exact 95% CI for  $\sigma^2$ , and, from this, an exact 95% CI for  $\gamma$ . Calculate the interval and compare with the approximate CI developed in **H**.

## Exercise 2

Exercise Rice 8: 8 (a) and (b) only.

**Hint for (a):** Use mle.

**Hint for (b):** Use section 8.5.3 – in particular the last paragraph before example **B**.  
(Notice that there is a hidden application of Slutsky's lemma in Rice's argument).